

# Package ‘sensitivymv’

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**Type** Package

**Title** Sensitivity Analysis in Observational Studies

**Version** 1.4.4

**Maintainer** Paul R. Rosenbaum <rosenbaum@wharton.upenn.edu>

**Description** The package performs a sensitivity analysis in an observational study using an M-statistic, for instance, the mean. The main function in the package is `senmv()`, but `amplify()` and `truncatedP()` are also useful. The method is developed in Rosenbaum Biometrics, 2007, 63, 456-464, <[doi:10.1111/j.1541-0420.2006.00717.x](https://doi.org/10.1111/j.1541-0420.2006.00717.x)>.

**License** GPL-2

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**LazyData** true

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sensitivitymv-package *Sensitivity Analysis in Observational Studies*

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## Description

The package performs a sensitivity analysis in an observational study using an M-statistic, for instance, the mean. The main function in the package is `senmv()`, but `amplify()` and `truncatedP()` are also useful. The method is developed in Rosenbaum Biometrics, 2007, 63, 456-464, <doi:10.1111/j.1541-0420.2006.00717.x>.

## Details

The DESCRIPTION file:

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Package:      sensitivitymv
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Title:        Sensitivity Analysis in Observational Studies
Version:      1.4.4
Authors@R:   person(given = c("Paul", "R."), family = "Rosenbaum", role = c("aut", "cre"), email = "rosenbaum@wharton.upenn.edu")
Maintainer:  Paul R. Rosenbaum <rosenbaum@wharton.upenn.edu>
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```

Index of help topics:

```
amplify      Amplification of sensitivity analysis in
              observational studies.
erpcp        DNA Damage Among Welders
lead150      Smoking and lead in 150 matched 1-5 sets.
lead250      Smoking and lead in 250 matched pairs.
mercury      NHANES Mercury/Fish Data
mscorev      Computes the M-scores used by senmv.
mtm          DNA damage from exposure to chromium
multrnks     Approximate scores for ranks.
newurks     Approximate scores for ranks of row ranges.
senmv        Sensitivity analysis in observational studies
              using Huber-Maritz M-statistics.
```

sensitivitymv-package	Sensitivity Analysis in Observational Studies
separable1v	Asymptotic separable calculations internal to other functions.
tbmetaphase	Genetic damage from drugs used to treat TB
truncatedP	Truncated product of P-values.
truncatedPbg	Truncated product of P-values using the mixture formula.

**Author(s)**

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**References**

- Rosenbaum, P. R. (2007) <doi:10.1111/j.1541-0420.2006.00717.x> Sensitivity analysis for m-estimates, tests and confidence intervals in matched observational studies. *Biometrics*, 2007, 63, 456-464.
- Rosenbaum, P. R. (2010) <doi:10.1007/978-1-4419-1213-8> Design of Observational Studies. New York: Springer Section 2.9 explains randomization inference with M-statistics in an example with 5 matched pairs.
- Rosenbaum, P. R. (2013) <doi:10.1111/j.1541-0420.2012.01821.x> Impact of multiple matched controls on design sensitivity in observational studies. *Biometrics*, 2013, 69, 118-127.
- Rosenbaum, P. R. (2014) <doi:10.1080/01621459.2013.879261> Weighted M-statistics with superior design sensitivity in matched observational studies with multiple controls. *Journal of the American Statistical Association*, 109(507), 1145-1158.
- Rosenbaum, P. R. (2015). Two R packages for sensitivity analysis in observational studies. *Observational Studies*, 1(1), 1-17. Free on-line at [obsstudies.org](http://obsstudies.org)
- Rosenbaum, P. R. (2017) *Observation and Experiment: An Introduction to Causal Inference*. Cambridge, MA: Harvard University Press. Chapter 9 discusses sensitivity analysis.

**Examples**

```
# Example reproduces parts of sect. 4.3 in Rosenbaum (2007)
data(tbmetaphase)
senmv(tbmetaphase, gamma=2, trim=1)
```

---

amplify

*Amplification of sensitivity analysis in observational studies.*

---

**Description**

Use of this function is illustrated in Rosenbaum (2015). Uses the method in Rosenbaum and Silber (2009) to interpret a value of the sensitivity parameter gamma, for instance the parameter in the senmv function. Each value of gamma amplifies to a curve (lambda,delta) in a two-dimensional sensitivity analysis, the inference being the same for all points on the curve. That is, a one-dimensional sensitivity analysis in terms of gamma has a two-dimensional interpretation.

**Usage**

```
amplify(gamma, lambda)
```

**Arguments**

gamma	gamma > 1 is the value of the sensitivity parameter, for instance the parameter in <code>senmv.length(gamma)&gt;1</code> will generate an error.
lambda	lambda is a vector of values > gamma. An error will result unless <code>lambda[i] &gt; gamma &gt; 1</code> for every i.

**Details**

A single value of gamma, say  $\gamma = 3.5$  in the example, corresponds to a curve of values of  $(\lambda, \delta)$ , including (4, 26), (6,8), (8,6), and (11,5) in the example. An unobserved covariate that is associated with a  $\lambda = 6$  fold increase in the odds of treatment and a  $\delta = 8$  fold increase in the odds of a positive pair difference is equivalent to  $\gamma = 3.5$ .

The curve is  $\gamma = (\lambda \cdot \delta + 1) / (\lambda + \delta)$ . Amplify is given one gamma and a vector of lambdas and solves for the vector of deltas. The calculation is elementary.

This interpretation of gamma is developed in detail in Rosenbaum and Silber (2009), and it makes use of Wolfe's (1974) family of semiparametric deformations of an arbitrary symmetric distribution. Strictly speaking, the amplification describes matched pairs, not matched sets. The `senmv` function views a k-to-1 matched set with k controls matched to one treated individual as a collection of k correlated treated-minus-control matched pair differences; see Rosenbaum (2007). For matched sets, it is natural to think of the amplification as describing any one of the k matched pair differences in a k-to-1 matched set. The curve has asymptotes that the function `amplify` does not compute: gamma corresponds with  $(\lambda, \delta) = (\gamma, \infty)$  and  $(\infty, \gamma)$ . A related though distinct idea is developed in Gastwirth et al (1998). The two approaches agree when the outcome is binary, that is, for McNemar's test.

**Value**

Returns a vector of values of delta of `length(lambda)` with names lambda.

**Note**

The example expands the discussion of Table 1 in Rosenbaum (2007). The study is insensitive to a bias of  $\gamma = 3.5$ . An unobserved covariate associated with a  $\lambda = 6$  fold increase in the odds of treatment and a  $\delta = 8$  fold increase in the odds of positive pair difference is equivalent to  $\gamma = 3.5$ . Also,  $\gamma = 3.5$  is equivalent to  $(\lambda, \delta) = (4, 26)$ , (8,6) and (11,5).

**Author(s)**

Paul R. Rosenbaum

## References

- Gastwirth, J. L., Krieger, A. M., Rosenbaum, P. R. (1998) <doi:10.1093/biomet/85.4.907> Dual and simultaneous sensitivity analysis for matched pairs. *Biometrika*, 85, 907-920.
- Rosenbaum, P. R. (2007) <doi:10.1111/j.1541-0420.2006.00717.x> Sensitivity analysis for m-estimates, tests and confidence intervals in matched observational studies. *Biometrics*, 2007, 63, 456-464.
- Rosenbaum, P. R. (2015). Two R packages for sensitivity analysis in observational studies. *Observational Studies*, 1(1), 1-17. Free on-line at [obsstudies.org](http://obsstudies.org).
- Rosenbaum, P. R. and Silber, J. H. (2009) <doi:10.1198/jasa.2009.tm08470> Amplification of sensitivity analysis in observational studies. *Journal of the American Statistical Association*, 104, 1398-1405.
- Rosenbaum, P. R. (2017) *Observation and Experiment: An Introduction to Causal Inference*. Cambridge, MA: Harvard University Press. Amplification is discussed in Chapter 9.
- Rosenbaum, P. R. (2020) <doi:10.1007/978-3-030-46405-9> *Design of Observational Studies*, 2nd edition. New York: Springer. Section 3.6 discusses amplification.
- Wolfe, D. A. (1974) A characterization of population weighted symmetry and related results. *Journal of the American Statistical Association*, 69, 819-822.

## Examples

```
data(erpcp)
senmv(erpcp, gamma=3.5, trim=1)
amplify(3.5, 6)
amplify(3.5, c(4, 6, 8, 11))
```

---

erpcp

*DNA Damage Among Welders*

---

## Description

Matched pairs of a welder and a control, matching for age and smoking. The values are DNA elution rates through polycarbonate filters with proteinase K (or erpcp). Data are originally from Werfel et al. (1998) and were used as an example in Rosenbaum (2007). Data are used to illustrate the `senmv` function in the `sensitivitymv` package.

## Usage

```
data(erpcp)
```

## Format

A data frame with 39 observations on the following 2 variables.

welder erpcp value for the welder

control erpcp value for the matched control

**Source**

Werfel et al. (1998).

**References**

Rosenbaum, P. R. Sensitivity analysis for m-estimates, tests and confidence intervals in matched observational studies. *Biometrics*, 2007, 63, 456-464.

Werful, U., Langen, V., Eickhoff, I. et al. Elevated DNA strand breakage frequencies in lymphocytes of welders exposed to chromium and nickel. *Carcinogenesis*, 1998, 19, 413-418.

**Examples**

```
data(erpcp)
```

---

lead150	<i>Smoking and lead in 150 matched 1-5 sets.</i>
---------	--

---

**Description**

Comparison of 150 daily smokers and 5 matched control nonsmokers from NHANES in terms of blood lead levels in  $\mu\text{-g/L}$ . The matching controlled for gender, age, race education level, and household income. The data were used in Rosenbaum (2013) to contrast design sensitivity with matched pairs and with multiple controls. See also lead250. Used as an example to illustrate the `senmv` function in the `sensitivitymv` package.

**Usage**

```
data(lead150)
```

**Format**

Column 1 is the lead level for the daily smoker. Columns 2 through 6 are lead levels for the five matched control nonsmokers.

**Details**

See the appendix to Rosenbaum (2013).

**Source**

NHANES 2007-2008 files from CDC and Rosenbaum (2013). Under a conventional Gaussian model, 250 pairs and 150 1-5 matched sets should have the same standard error; however, the design sensitivity is not the same.

**References**

Rosenbaum, P. R. (2013) Impact of multiple matched controls on design sensitivity in observational studies. *Biometrics*, 2013, 69, 118-127.

**Examples**

```
data(lead150)
```

---

lead250	<i>Smoking and lead in 250 matched pairs.</i>
---------	---

---

**Description**

Comparison of 250 daily smokers and one matched control nonsmoker from NHANES in terms of blood lead levels in  $\mu\text{-g/L}$ . The matching controlled for gender, age, race education level, and household income. The data were used in Rosenbaum (2013) to contrast design sensitivity with matched pairs and with multiple controls. See also lead150. Used as an example to illustrate the `senmv` function in the `sensitivitymv` package.

**Usage**

```
data(lead250)
```

**Format**

Column 1 describes the smoker and column 2 describes the matched nonsmoker.

**Details**

See the appendix to Rosenbaum (2013). Under a conventional Gaussian model, 250 pairs and 150 1-5 matched sets should have the same standard error; however, the design sensitivity is not the same.

**Source**

NHANES 2007-2008 from CDC and Rosenbaum (2013).

**References**

Rosenbaum, P. R. (2013) Impact of multiple matched controls on design sensitivity in observational studies. *Biometrics*, 2013, 69, 118-127.

**Examples**

```
data(lead250)
```

---

mercury

*NHANES Mercury/Fish Data*

---

### **Description**

Data from NHANES. Matrix 397 x 3. n=397 treated people who ate at least 15 servings of fish or shellfish during the previous month are matched to two controls who ate at most one serving of fish or shellfish. The values in mercury record the level of methylmercury in blood in  $\mu\text{-g/dl}$ . Column 1 is treated, columns 2 and 3 are controls. Specifically, column 2 describes an individual who ate no fish or shellfish in the previous month, while column 3 describes an individual who ate exactly one serving of fish or shellfish. The data were used as an example in Rosenbaum (2014).

### **Usage**

```
data(mercury)
```

### **Format**

A data frame with 397 observations on the following 3 variables.

Treated Mercury level for treated individual.

Zero Mercury level for a control who ate no fish/shellfish

One Mercury level for a control who ate one serving of fish/shellfish

### **Source**

Data is originally from NHANES 2009-2010.

### **References**

Rosenbaum, P. R. (2014). Weighted M-statistics with superior design sensitivity in matched observational studies with multiple controls. *Journal of the American Statistical Association*, 109(507), 1145-1158. <doi:10.1080/01621459.2013.879261>

### **Examples**

```
data(mercury)
```



---

mscorev

*Computes the M-scores used by senmv.*


---

**Description**

Computes the M-scores used by senmv for sensitivity analysis in observational studies using Huber-Maritz M-tests.

**Usage**

```
mscorev(ymat, inner = 0, trim = 2.5, qu = 0.5, TonT = FALSE)
```

**Arguments**

ymat	ymat is a matrix as described in the documentation for senmv.
inner	inner is the parameter described in the documentation for senmv.
trim	trim is the parameter described in the documentation for senmv.
qu	qu is the lambda parameter described in the documentation for senmv.
TonT	If TonT=FALSE, then the total score in set (row) i is divided by the number ni of individuals in row i, as in expression (8) in Rosenbaum (2007). If TonT=TRUE, then the division is by ni-1, not by ni, and there is a further division by the total number of matched sets. See the discussion of TonT in the documentation for senmv.

**Value**

Generally, a matrix with the same dimensions as ymat containing the M-scores. Exception: if a matched set does not contain at least one treated subject and at least one control, then that set will not appear in the result, and the result will have fewer rows than ymat. However, if a matched set has several controls but no treated subject, then these controls will contribute to the estimate of the scale parameter, typically the median absolute pair difference.

**Note**

The example reproduces Table 3 in Rosenbaum (2007).

**Author(s)**

Paul R. Rosenbaum

**References**

Rosenbaum, P. R. (2007) <doi:10.1111/j.1541-0420.2006.00717.x> Sensitivity analysis for m-estimates, tests and confidence intervals in matched observational studies. *Biometrics*, 2007, 63, 456-464.

Rosenbaum, P. R. (2015). Two R packages for sensitivity analysis in observational studies. *Observational Studies*, 1(1), 1-17. Free on-line at [obsstudies.org](http://obsstudies.org)

**Examples**

```
data(tbmetaphase)
mscorev(tbmetaphase, trim=1)
```

---

mtm

*DNA damage from exposure to chromium*

---

**Description**

The data are from a study by Meibian et al. (2008) concerning possible damage to human DNA from occupational exposure to chromium. There were three matched groups, the control group (cmtm), a low exposure group (e2mtm) and a high exposure group (e1mtm). The exposed individuals all worked at a tannery where chromium was used in the tanning of leather. The highly exposed group (e1) worked at tanning leather. The low exposure group (e2) worked at the same tannery but did not tan leather. The reported values are the mean tail moment (mtm) of the comet assay, a measure of damage to DNA. High values of mtm indicate greater damage.

**Usage**

```
data(mtm)
```

**Format**

Each row of mtm is a matched set. The columns refer to the treatment groups mentioned in the description.

**Details**

These data were used as an example of approximate evidence factors in Rosenbaum (2011). Under the null hypothesis  $H_0$  of no treatment effect, there are two approximately independent tests of  $H_0$  subject to different biases of nonrandom selection, specifically the comparison of 30 controls and 60 matched tannery workers, and the comparison of 30 low and 30 high exposure tannery workers. The two comparisons may be subjected to sensitivity analyses, say using `senmv`, and the results of these two analyses may be combined, for instance using Fisher's method of combining independent P-values. See the documentation for `truncatedP` or `truncatedPbg` for an example.

**Source**

Meibian et al. (2008). Used as an example in Rosenbaum (2011).

**References**

Meibian, Z., Zhijian, C., Qing, C. et al. (2008) Investigating DNA damage in tannery workers occupationally exposed to tivalent chromium using the comet assay. *Mutation Research* 654, 45-51.

Rosenbaum, P. R. (2011) Some approximate evidence factors in observational studies. *Journal of the American Statistical Association*, 2011, 106, 285-295. <doi:10.1198/jasa.2011.tm10422>

Rosenbaum, P. R. (2015). Two R packages for sensitivity analysis in observational studies. *Observational Studies*, 1(1), 1-17. Free on-line at [obsstudies.org](http://obsstudies.org)

### Examples

```
data(mtm)
```

---

multrnks	<i>Approximate scores for ranks.</i>
----------	--------------------------------------

---

### Description

Of limited interest to most users, this function is sometimes called by `senmv`. The function calculates the large sample approximation to a rank score transformation in Lemma 1, expression (9) of Rosenbaum (2011). See also Rosenbaum (2014) and the `sensitivymw` package.

### Usage

```
multrnks(rk, m1 = 2, m2 = 2, m = 2)
```

### Arguments

<code>rk</code>	A vector of ranks that may include average ranks for ties.
<code>m1</code>	One of three rank score parameters in Rosenbaum (2011), specifically $m_1 = \underline{m}$ .
<code>m2</code>	One of three rank score parameters in Rosenbaum (2011), specifically $m_2 = \overline{m}$ .
<code>m</code>	One of three rank score parameters in Rosenbaum (2011), specifically $m = m$ .

### Value

Vector of length(`rk`) containing the scores for the ranks in `rk`.

### Author(s)

Paul R. Rosenbaum

### References

Rosenbaum, P. R. (2011) <doi:10.1111/j.1541-0420.2010.01535.x> A new u-statistic with superior design sensitivity in matched observational studies. *Biometrics* 67, 1017-1027.

Rosenbaum, P. R. (2014) <doi:10.1080/01621459.2013.879261> Weighted M-statistics with superior design sensitivity in matched observational studies with multiple controls. *Journal of the American Statistical Association*, 109(507), 1145-1158.

Rosenbaum, P. R. (2024) <doi:10.1080/01621459.2023.2221402> Bahadur efficiency of observational block designs. *Journal of the American Statistical Association*, 119(547), 1871-1881.

**Examples**

```
multrnks(1:10)
multrnks(1:10,m1=12,m2=20,m=20)
```

---

newurks	<i>Approximate scores for ranks of row ranges.</i>
---------	--

---

**Description**

Of limited interest to most users, this function is sometimes called by `senmv`. The function calculates the ranges for each row of `smat`, scores their ranks using the large sample approximation to a rank score transformation in Lemma 1, expression (9) of Rosenbaum (2011), as implemented in the function `multrnks`, and multiplies the rows by these rank scores.

**Usage**

```
newurks(smat, m1 = 2, m2 = 2, m = 2)
```

**Arguments**

<code>smat</code>	A matrix. When called by <code>senmv</code> , <code>smat</code> is a matrix produced by <code>mscorev</code> .
<code>m1</code>	One of three rank score parameters in Rosenbaum (2011), specifically $m1 = \underline{\text{underline}}(m)$ .
<code>m2</code>	One of three rank score parameters in Rosenbaum (2011), specifically $m2 = \overline{\text{overline}}(m)$ .
<code>m</code>	One of three rank score parameters in Rosenbaum (2011), specifically $m = m$ .

**Value**

The function calculates the ranges for each row of `smat`, scores the ranks of the ranges using the large sample approximation to a rank score transformation in Lemma 1, expression (9) of Rosenbaum (2011), as implemented in the function `multrnks`, and multiplies the rows of `smat` by these rank scores, returning the result.

**Author(s)**

Paul R. Rosenbaum

**References**

Rosenbaum, P. R. (2011) <doi:10.1111/j.1541-0420.2010.01535.x> A new u-statistic with superior design sensitivity in matched observational studies. *Biometrics* 67, 1017-1027.

Rosenbaum, P. R. (2014) <doi:10.1080/01621459.2013.879261> Weighted M-statistics with superior design sensitivity in matched observational studies with multiple controls. *Journal of the American Statistical Association*, 109(507), 1145-1158.

Rosenbaum, P. R. (2024) <doi:10.1080/01621459.2023.2221402> Bahadur efficiency of observational block designs. *Journal of the American Statistical Association*, 119(547), 1871-1881.

**Examples**

```
data(lead150)
head(newurks(mscorev(lead150)))
```

---

senmv	<i>Sensitivity analysis in observational studies using Huber-Maritz M-statistics.</i>
-------	---

---

**Description**

Computes the large sample approximation to the upper bound on the one sided P-value testing the null hypothesis of no treatment effect in a matched observational study with one or more controls matched to each treated subject. Uses Huber's M-statistics as test statistics in the sense proposed by Maritz (1979). Use of `senmv()` is discussed and illustrated in Rosenbaum (2015a). As a possible alternative to `senmv()`, please consider the functions `senm()` and `senmCI()` functions in the `sensitivitymult` package – they are more recent implementations that add confidence intervals with simpler input.

**Usage**

```
senmv(y, gamma = 1, method = NULL, inner = 0, trim = 2.5,
      lambda = 1/2, tau = 0, TonT=FALSE)
```

**Arguments**

y	If y is a vector, then y is the vector of treated-minus-control pair differences in outcomes in $n=\text{length}(y)$ matched pairs. If y is an n by J matrix, then: (i) the rows are n matched sets, (ii) the first column is the treated response in a set, columns 2 to J contain the responses of controls in the same matched set. If a matched set has fewer than J-1 controls, then NA's fill the empty spaces in y. The test is one-sided, testing no treatment effect against the alternative hypothesis that treated responses are higher than control responses. Replace y by -y to test against the alternative that treated responses are lower than control responses; see the <code>mtm</code> example.
gamma	gamma is the sensitivity parameter, $\text{gamma}=1$ for a randomization test, $\text{gamma}>1$ for sensitivity bounds. Use of $\text{gamma}<1$ will generate an error. This parameter gamma is denoted by the upper case Greek letter gamma in the cited literature, for instance Rosenbaum (2007).
method	If method is NULL, then the method is determined by the parameters, namely inner, trim, lambda, and TonT. If method is not NULL, then these parameters are set according to the selected method and stated values of the parameters are ignored. The default values of the parameters are equivalent to <code>method="h"</code> . (i) <code>method = "h"</code> (Huber) is the default and sets <code>inner=0</code> , <code>trim=2.5</code> , <code>lambda = 1/2</code> , <code>TonT=FALSE</code> . Its psi function levels off at 2.5 times the median ( $\text{lambda} = 1/2$ ) of the absolute pair differences. Method h is often a good choice in small samples with few pairs or sets. Method h is often a good choice when the number of

controls in each matched set is 6 or more. For long tailed distributions, it may be better to use the same parameter values except with a lower value of trim, perhaps  $\text{trim} = 2$  or  $\text{trim} = 1$ . (ii)  $\text{method} = "i"$  (inner) often has better design sensitivity than  $\text{method} = "h"$ . It has  $\text{inner} = 1/2$ , ignoring absolute pair differences smaller than half the median. It has  $\text{inner}=1/2$ ,  $\text{trim}=2.5$ ,  $\text{lambda} = 1/2$ ,  $\text{TonT}=\text{FALSE}$ . Method  $i$  is often a good choice for a moderate to large sample of matched pairs, say 100 or more pairs. Method  $i$  is often a good choice for a moderate to large number of matched triples, say 100 or more. See Rosenbaum (2013) for comparisons of methods  $h$  and  $i$ . (iii)  $\text{method} = "t"$  (permutational t-test) permutes the observations themselves, not scores derived from the observations. It has  $\text{inner}=0$ ,  $\text{trim}=\text{Inf}$ ,  $\text{lambda} = 1/2$ ,  $\text{TonT}=\text{TRUE}$ . This is the only method that sets  $\text{TonT}=\text{TRUE}$ . The test statistic gives the mean over sets of the mean treated-minus-control difference within sets, that is, the usual estimate of the effect of the treatment on the treated. See the example for mercury using  $\text{method}="t"$  and the discussion of the  $\text{TonT}$  parameter.

inner	Inner trimming to increase design sensitivity. See the discussion of lambda. Use of $\text{inner}<0$ or $\text{inner}>\text{trim}$ will generate an error.
trim	Outer trimming for resistance to outliers. Setting $\text{trim} = \text{Inf}$ does no trimming. See the discussion of lambda.
lambda	Observations are scaled by the lambda quantile of the absolute pair differences, defaulting to the median of all paired absolute differences; see Rosenbaum (2007, section 4.2) for a precise definition in the case of multiple controls. If the lambda quantile of the absolute pair differences is 0, then scaling by 0 is impossible and an error may result; the solution is to increase lambda. If $q_u$ is the lambda quantile, absolute pair differences smaller than $\text{inner}*q_u$ receive weight 0, absolute pair differences larger than $q_u*\text{trim}$ receive weight 1, and between $\text{inner}*q_u$ and $\text{trim}*q_u$ weights increase linearly from 0 to 1. Use of $\text{lambda}<=0$ or $\text{lambda}>=1$ will generate an error. If $\text{inner}=0$ and $\text{trim}=\text{Inf}$ , then this results in the permutational t-test in which the observations themselves are permuted, and in this case lambda is not used. Taking $\text{lambda} = .95$ and $\text{trim} = 1$ is similar to trimming 5 percent of the pair differences.
tau	If $\text{tau}=0$ , senmv tests the null hypothesis of no treatment effect. If tau is not 0, senmv tests the null hypothesis that the treatment effect is an additive shift of tau against the alternative that the effect is larger than tau.
TonT	$\text{TonT}$ refers to the effect of the treatment on the treated. The default is $\text{TonT}=\text{FALSE}$ . For all methods except $\text{method}="t"$ , $\text{TonT}=\text{FALSE}$ . $\text{TonT}=\text{TRUE}$ is mostly relevant to the permutational t-test, $\text{method}="t"$ , but it remains an option whenever $\text{is.null}(\text{method})=\text{TRUE}$ . If $\text{TonT}=\text{FALSE}$ , then the total score in set (row) $i$ is divided by the number $n_i$ of individuals in row $i$ , as in expression (8) in Rosenbaum (2007). This division by $n_i$ has few consequences when every matched set has the same number of individuals, but when set sizes vary, dividing by $n_i$ is intended to increase efficiency by weighting inversely as the variance; see the discussion in section 4.2 of Rosenbaum (2007). If $\text{TonT}=\text{TRUE}$ , then the division is by $n_i-1$ , not by $n_i$ , and there is a further division by the total number of matched sets to make it a type of mean. If $\text{TonT}=\text{TRUE}$ , then the statistic is the mean over matched sets of the treated-score minus the mean-of-the-control-scores within matched sets, so it is weighted to estimate the effect of the treatment on the treated. See the mercury example with $\text{method}="t"$ .

## Details

The one-sided alternative hypothesis is that treatment increases the level of response. The `senmv` function has as its default the use of Huber's (1981) psi function, which is similar in certain respects to a trimmed mean. By default, the trimming occurs (i.e., Huber's psi function becomes horizontal) at  $\text{trim}=2.5$  times the median (i.e.,  $\lambda = 1/2$ ) of the absolute pair differences; see Maritz (1979) for the paired case and see Rosenbaum (2007, section 4) for matched sets with one or more controls. With  $\text{trim} = 2.5$ , Huber's psi function is somewhat analogous to a trimmed mean that trims five percent from each tail. If  $\text{trim}$  is changed to 1, then the psi function is horizontal at 1 times the median (i.e.,  $\lambda=1/2$ ) of the absolute pair differences and is somewhat analogous to a midmean. `method="t"` performs the permutational t-test, the permutation distribution of a type of mean; see, for instance, Pitman (1937).

The `senmv` function can be used in the calculation of approximate evidence factors, as discussed and illustrated in the documentation for the `truncatedP` and `truncatedPbg` functions.

## Value

<code>pval</code>	Approximate upper bound on the one-sided P-value.
<code>deviate</code>	Deviate that is compared to the upper tail of the standard Normal distribution to obtain the P-value.
<code>statistic</code>	Value of the test statistic.
<code>expectation</code>	Maximum null expectation of the test statistic for the given value of $\gamma$ .
<code>variance</code>	Among null distributions that yield the maximum expectation, variance is the maximum possible variance for the given value of $\gamma$ . See Rosenbaum (2007, Section 4) and Gastwirth, Krieger and Rosenbaum (2000).

## Note

Example `erpcp` reproduces parts of Table 1 in Rosenbaum (2007), example `tbmetaphase` reproduces parts of section 4.3 in Rosenbaum (2007). Examples `lead150` and `lead250` reproduce parts of Table 1 of Rosenbaum (2013).

Example `mtm` almost reproduces parts of section 6 of Rosenbaum (2011). The second P-value bound differs slightly from that reported in the paper. The reason is that the paper used one scale factor, the median absolute pair difference, for both tests based on all pairs, whereas `senmv` applied to the last two columns of `mtm` uses a scale factor derived just from these two columns. There is no obvious reason to prefer one approach over the other, and the approach in the paper would require a separate R function. Note that in this example, `y` is `-mtm` because we test that the controls are low rather than high.

See the help file for `mscorev` to reproduce Table 3 in Rosenbaum (2007).

The mercury example illustrates the `TonT` parameter in a trivial case, because every treated subject has two controls. Note that the test statistic is a kind of mean when `TonT=TRUE`, but is merely a sum when `TonT=FALSE`. The weighting used when `TonT=FALSE` means the test statistic uses information efficiently under certain models, but the test statistic itself is not an estimate of anything. Note the effect of setting `tau = 2` is to reduce the test statistic by 2 when `method="t"`.

## Author(s)

Paul R. Rosenbaum

## References

- Main references: Rosenbaum, P. R. (2007) <doi:10.1111/j.1541-0420.2006.00717.x> Sensitivity analysis for m-estimates, tests and confidence intervals in matched observational studies. *Biometrics*, 2007, 63, 456-464. This paper is the main reference for the `senmv` function when weights are not used,  $m_1=m_2=m=1$ .
- Rosenbaum, P. R. (2013) <doi:10.1111/j.1541-0420.2012.01821.x> Impact of multiple matched controls on design sensitivity in observational studies. *Biometrics*, 2013, 69, 118-127. Evaluates the performance of the methods in the paper above, and in particular provides a basis for selecting the parameter values for `senmv`. In particular, this paper compares methods `h`, `i`, and `t`.
- Rosenbaum, P. R. (2014) <doi:10.1080/01621459.2013.879261> Weighted M-statistics with superior design sensitivity in matched observational studies with multiple controls. *Journal of the American Statistical Association*, 109(507), 1145-1158.
- Rosenbaum, P. R. (2015a). Two R packages for sensitivity analysis in observational studies. *Observational Studies*, 1(1), 1-17. Free on-line at [obsstudies.org](http://obsstudies.org) ————— Additional references: Cox, D. R. and Reid, N. *Theory of the Design of Experiments*. New York: Chapman and Hall/CRC. Chapter 2 discusses randomization inference, in particular an unmatched version of the permutation distribution of the treated minus control difference in mean responses, or the two-sample (unmatched) permutational t-test. Although there is a large old literature on tests that permute the observations, this recent discussion is written in a modern style.
- Fisher, R. A. (1935) *Design of Experiments*. Edinburgh: Oliver and Boyd. Chapter 3 contains an early, conceptual discussion of the permutation distribution of the mean or the permutational t-test.
- Huber, P. (1981) *Robust Statistics*. New York: Wiley, 1981. Huber first proposed the use of m-statistics in 1964 in a paper in the *Annals*.
- Maritz, J. S. (1979) Exact robust confidence intervals for location. *Biometrika* 1979, 66, 163-166. Proposed exact permutation tests using m-statistics that Maritz inverts to obtain exact confidence limits. The subtle aspect is the scaling which must be invariant to treatment assignment under the null hypothesis, so it differs from the scaling used by Huber.
- Gastwirth, J. L., Krieger, A. M., and Rosenbaum, P. R. (2000) Asymptotic separability in sensitivity analysis. *Journal of the Royal Statistical Society B* 2000, 62, 545-556. Provides a general large sample approximation when matching with multiple controls, as used in Rosenbaum (2007, Section 4).
- Pitman, E. J. G. (1937) Significance tests which may be applied to samples from any populations. *JRSS-supplement* (later called series B), 4, 119-130. An early technical discussion of the permutation distribution of the sample mean, or the permutational t-test.
- Rosenbaum, P. R. (2010) *Design of Observational Studies*. New York: Springer 2010. <doi:10.1111/j.1541-0420.2012.01821.x> Section 2.9 contains an elementary textbook discussion of Maritz's permutation distribution for m-statistics.
- Rosenbaum, P. R. (2025) *An Introduction to the Theory of Observational Studies*. New York: Springer. ISBN 978-3-031-90493-6.
- Rosenbaum, P. R. (2011) Some approximate evidence factors in observational studies. *Journal of the American Statistical Association*, 2011, 106, 285-295. <doi:10.1198/jasa.2011.tm10422> The method described in this paper may be implemented using `senmv`. To do this, one uses `senmv` several times, combining the resulting one-sided P-value bounds, perhaps using Fisher's method for combining P-values. In the example of this paper,  $y$  is  $n \times 3$  for three groups in matched triples, and



the paper uses Fisher's method to combine the P-value bounds from `senmv(y)` and `senmv(y[,2:3])` for an appropriately defined `y`. See the `mtm` example in the documentation for `truncatedP` and `truncatedPbg`.

Rosenbaum, P. R. (2015b). How to see more in observational studies: Some new quasi-experimental devices. *Annual Review of Statistics and Its Application*, 2, 21-48. Discusses evidence factors. <doi/10.1146/annurev-statistics-010814-020201>

Rosenbaum, P. R. (2017) *Observation and Experiment: An Introduction to Causal Inference*. Cambridge, MA: Harvard University Press. Sensitivity analysis is discussed in Chapter 9.

Welch, B. L. (1937) On the z-test in randomized blocks. *Biometrika* 29, 21-52. As in Pitman (1937) above, discusses permutation inference in which the responses are permuted, essentially a permutational F-test. Expresses causal effects as comparisons of potential responses under alternative treatments.

## Examples

```
# This example reproduces parts of Table 1 in Rosenbaum (2007).
data(erpcp)
senmv(erpcp,gamma=3,trim=1)
senmv(erpcp,gamma=2,trim=1)
senmv(erpcp,gamma=2,trim=1,tau=0.34)
senmv(erpcp,gamma=2,trim=1,tau=0.18)
senmv(erpcp,gamma=2,trim=1,tau=0.185)

# Example reproduces parts of sect. 4.3 in Rosenbaum (2007)
data(tbmetaphase)
senmv(tbmetaphase,gamma=2,trim=1)
senmv(tbmetaphase,gamma=2,trim=1,tau=0.94)
senmv(tbmetaphase,gamma=2,trim=1,tau=0.945)

# Example reproduces part of Table 1 in Rosenbaum (2013)
data(lead150)
senmv(lead150,gamma=2,trim=2)
data(lead250)
senmv(lead250,gamma=2,trim=2)

# Example reproduces parts of of Rosenbaum (2011). See documentation for truncatedP.
data(mtm)
senmv(-mtm,gamma=11.7,trim=1)
senmv(-mtm[,2:3],gamma=2.1,trim=1)

# Illustrates method = "i"
data(mercury)
senmv(mercury,gamma=15)
senmv(mercury,gamma=15,method="i")

# Illustrates TonT=TRUE as in method="t". See note above.
data(mercury)
mean(mercury[,1])-(mean(mercury[,2])+mean(mercury[,3]))/2
senmv(mercury,gamma=15,trim=Inf,TonT=TRUE)$statistic
senmv(mercury,gamma=15,method="t")$statistic
senmv(mercury,gamma=15,method="t",tau=1)$statistic
```

```
senmv(mercury,gamma=15,method="t",tau=2)$statistic
senmv(mercury,gamma=15,trim=Inf,TonT=FALSE)$statistic
```

---

separable1v

*Asymptotic separable calculations internal to other functions.*


---

## Description

This general purpose function is internal to other functions, such as `senmv` in the `sensitivymv` package. The function performs the asymptotic separable calculations described in Gastwirth, Krieger and Rosenbaum (2000), as used in section 4 of Rosenbaum (2007). The example is equivalent to `senmv(lead150,gamma=2,method="t")`.

## Usage

```
separable1v(ymat, gamma = 1)
```

## Arguments

<code>ymat</code>	<code>ymat</code> is a matrix whose rows are matched sets and whose columns are matched individuals. The first column describes treated individuals. Other columns describe controls. If matched sets contain variable numbers of controls, NAs fill in empty spaces in <code>ymat</code> ; see the documentation for <code>senmv</code> . In <code>senmv</code> , the matrix <code>ymat</code> is created by <code>mscorev</code> . Instead, if there were no NAs and ranks within rows were used in <code>ymat</code> , then <code>separable1v</code> would perform a sensitivity analysis for the stratified Wilcoxon two-sample test. Applied directly to data, it performs a sensitivity analysis for the permutational t-test.
<code>gamma</code>	<code>gamma</code> is the value of the sensitivity parameter; see the documentation for the <code>senmv</code> function in the <code>sensitivymv</code> package. One should use a value of <code>gamma</code> $\geq 1$ .

## Details

For discussion of asymptotic separability in sensitivity analysis, see Gastwirth, Krieger and Rosenbaum (2000), Rosenbaum (2007, Section 4) and Rosenbaum (2018). The function `separable1v()` permits matched sets with one treated individual and a variable number of controls. If every matched set has the same number of controls, then the function `separable1k()` in the `sensitivymw` package will be somewhat faster. For strata of arbitrary size and structure, see the `senstrat` package and Rosenbaum (2018).

## Value

<code>pval</code>	Approximate upper bound on the one-sided P-value.
<code>deviate</code>	Deviate that is compared to the upper tail of the standard Normal distribution to obtain the P-value.
<code>statistic</code>	Value of the test statistic.

expectation	Maximum null expectation of the test statistic for the given value of gamma.
variance	Among null distributions that yield the maximum expectation, variance is the maximum possible variance for the given value of gamma.

**Author(s)**

Paul R. Rosenbaum

**References**

Gastwirth, J. L., Krieger, A. M., and Rosenbaum, P. R. (2000) <doi:10.1111/1467-9868.00249> Asymptotic separability in sensitivity analysis. *Journal of the Royal Statistical Society B* 2000, 62, 545-556.

Rosenbaum, P. R. (2007) <doi:10.1111/j.1541-0420.2006.00717.x> Sensitivity analysis for m-estimates, tests and confidence intervals in matched observational studies. *Biometrics*, 2007, 63, 456-464.

Rosenbaum, P. R. (2018) <doi:10.1214/18-AOAS1153> Sensitivity analysis for stratified comparisons in an observational study of the effect of smoking on homocysteine levels. *The Annals of Applied Statistics*, 12(4), 2312-2334.

**Examples**

```
data(lead150)
separable1v(lead150, gamma=2)
```

---

tbmetaphase

*Genetic damage from drugs used to treat TB*

---

**Description**

This is a matched comparison of the effects of two drug sequences, namely HRZ and H2R2Z2, for the treatment of tuberculosis. HRZ is a higher dose sequence than H2R2Z2. The outcome is a measure of genetic damage, namely the frequency of aberrant metaphases two months after treatment. Individuals were matched for the frequency of aberrant metaphases before treatment. 15 individuals treated with HRZ are matched to 1 or 2 controls treated with H2R2Z2. Each row is one matched set. If a set is a pair, the third element in a row is NA. The data are originally from Rao, Gupta and Thomas (1991) and were used as an example in Rosenbaum (2007, Table 3). Data are used to illustrate the `senmv` function in the `sensitivitymv` package.

**Usage**

```
data(tbmetaphase)
```

**Format**

A data frame with 15 observations on the following 3 variables.

HRZ Aberrant metaphases for individual treated with HRZ.

H2R2Z2.1 Aberrant metaphases for first matched individual treated with H2R2Z2.

H2R2Z2.2 Aberrant metaphases for second matched individual treated with H2R2Z2. For matched pairs, this is NA.

**References**

Rao, V. V. N. G., Gupta, E. V. V and Thomas, I. M. Chromosomal aberrations in tuberculosis patients before and after treatment with short-term chemotherapy. *Mutation Research* 1991, 259, 13-19.

Rosenbaum, P. R. Sensitivity analysis for m-estimates, tests and confidence intervals in matched observational studies. *Biometrics*, 2007, 63, 456-464.

**Examples**

```
data(tbmetaphase)
```

---

truncatedP	<i>Truncated product of P-values.</i>
------------	---------------------------------------

---

**Description**

Zaykin et al. (2002) proposed combining  $L = \text{length}(p)$  independent P-values  $p$  by taking the product of the P-values that are no larger than a truncation point  $\text{trunc}$ , namely  $w = \prod(p^{(p \leq \text{trunc})})$ . Computes the one P-value for the combination  $w$ . Use of `truncatedP` with evidence factors is illustrated in Rosenbaum (2015).

**Usage**

```
truncatedP(p, trunc = 0.2)
```

**Arguments**

p	a vector of P-values that are either independent or stochastically larger than the uniform distribution on the cube (see Rosenbaum (2011, Definition 1))
trunc	truncation point. Computes the product of all P-values in $p$ that are no larger than $\text{trunc}$ . For $\text{trunc} = 1$ , performs Fisher's method of combining independent P-values.

## Details

Zaykin et al. (2002) proposed combining  $L = \text{length}(p)$  independent P-values  $p$  by taking the product of the P-values that are no larger than a truncation point  $\text{trunc}$ , namely  $w = \prod(p^{(p \leq \text{trunc})})$ . For  $\text{trunc} = 1$ , this is Fisher's method for combining independent P-values. The method also works for certain kinds of fairly inconsequential dependence; see Rosenbaum (2011, section 2).

`truncatedP` computes the one P-value for the combination  $w$  using the formula in Zaykin et al. (2002). The equivalent function `truncatedPbg` computes the exact same P-value for  $w$  using a binomial mixture of gamma distributions, as discussed by Hsu et al. (2013, section 3.1).

The truncated product or Fisher's method ( $\text{trunc} = 1$ ) may be used for sensitivity analyses with evidence factors; see Rosenbaum (2011) and the `mtm` example below. The truncated product with  $\text{trunc} < 1$  is useful in combining P-value upper bounds produced by sensitivity analyses, for instance those produced by `senmv`. These upper bounds eventually approach 1 for larger values of the sensitivity parameter, and  $\text{trunc} < 1$  eliminates these, often increasing power. See Hsu et al. (2013) for comparisons.

## Value

Returns the one P-value for the truncated product. For  $\text{trunc} < 1$ , the distribution is not continuous, but rather attaches positive probability to a P-value of 1.

## Note

Example `mtm` illustrates approximate evidence factors and almost reproduces parts of section 6 of Rosenbaum (2011) using the `mtm` data from Meibian et al. (2008). The second P-value bound differs slightly from that reported in the paper. The reason is that the paper used one scale factor, the median absolute pair difference, for both tests based on all pairs, whereas `senmv` applied to the last two columns of `mtm` uses a scale factor derived just from these two columns. There is no obvious reason to prefer one approach over the other, and the approach in the paper would require a separate R function. Note that in this example,  $y$  is `-mtm` because we test that the controls are low rather than high.

## Author(s)

Paul R. Rosenbaum

## References

Hsu, J. Y., Small, D. S. and Rosenbaum, P. R. (2013) <doi:10.1080/01621459.2012.742018> Effect modification and design sensitivity in observational studies. *Journal of the American Statistical Association*, 108, 135-148.

Meibian, Z., Zhijian, C., Qing, C. et al. (2008) Investigating DNA damage in tannery workers occupationally exposed to tivalent chromium using the comet assay. *Mutation Research* 654, 45-51.

Rosenbaum, P. R. (2010) <doi:10.1093/biomet/asq019> Evidence factors in observational studies. *Biometrika*, 97, 333-345.

Rosenbaum, P. R. (2011) <doi:10.1198/jasa.2011.tm10422> Some approximate evidence factors in observational studies. *Journal of the American Statistical Association*, 106, 285-295. <doi:10.1198/jasa.2011.tm10422>

Rosenbaum, P. R. (2015). Two R packages for sensitivity analysis in observational studies. *Observational Studies*, 1(1), 1-17. Free on-line at [obsstudies.org](http://obsstudies.org)

Rosenbaum, P. R. (2021) <doi:10.1201/9781003039648> *Replication and Evidence Factors in Observational Studies*. Chapman and Hall/CRC.

Zaykin, D. V., Zhivotovsky, L. A., Westfall, P. H. and Weir, B. S. (2002) Truncated product method of combining P-values. *Genetic Epidemiology*, 22, 170-185. <doi:10.1002/gepi.0042>

Zhang, K., Small, D. S., Lorch, S., Srinivas, S. and Rosenbaum, P. R. (2011) Using split samples and evidence factors in an observational study of neonatal outcomes. *Journal of the American Statistical Association*, 106, 511-524.

## Examples

```
# Evidence factor example: see note above.
data(mtm)
senmv(-mtm, gamma=11.7, trim=1)
senmv(-mtm[, 2:3], gamma=2.1, trim=1)
senmv(-mtm, gamma=12, trim=1)
senmv(-mtm[, 2:3], gamma=3, trim=1)
truncatedP(c(0.05167572, 0.1527849), trunc=1)
truncatedP(c(0.05167572, 0.1527849), trunc=.2)
```

---

truncatedPbg

*Truncated product of P-values using the mixture formula.*

---

## Description

Zaykin et al. (2002) proposed combining  $L = \text{length}(p)$  independent P-values  $p$  by taking the product of the P-values that are no larger than a truncation point  $\text{trunc}$ , namely  $w = \prod(p^{(p \leq \text{trunc})})$ . Computes the one P-value for the combination  $w$ .

## Usage

```
truncatedPbg(p, trunc = 0.2)
```

## Arguments

<code>p</code>	a vector of P-values that are either independent or stochastically larger than the uniform distribution on the cube (see Rosenbaum (2011, Definition 1))
<code>trunc</code>	truncation point. Computes the product of all P-values in <code>p</code> that are no larger than <code>trunc</code> . For <code>trunc = 1</code> , performs Fisher's method of combining independent P-values.

## Details

Zaykin et al. (2002) proposed combining  $L = \text{length}(p)$  independent P-values  $p$  by taking the product of the P-values that are no larger than a truncation point  $\text{trunc}$ , namely  $w = \prod(p^{(p \leq \text{trunc})})$ . For  $\text{trunc} = 1$ , this is Fisher's method for combining independent P-values. The method also works for certain kinds of fairly inconsequential dependence; see Rosenbaum (2011, section 2).

`truncatedP` computes the one P-value for the combination  $w$  using the formula in Zaykin et al. (2002). The equivalent function `truncatedPbg` computes the exact same P-value for  $w$  using a binomial mixture of gamma distributions, as discussed by Hsu et al. (2013, section 3.1).

The truncated product or Fisher's method ( $\text{trunc} = 1$ ) may be used for sensitivity analyses with evidence factors; see Rosenbaum (2011) and the `mtm` example below. The truncated product with  $\text{trunc} < 1$  is useful in combining P-value upper bounds produced by sensitivity analyses, for instance those produced by `senmv`. These upper bounds eventually approach 1 for larger values of the sensitivity parameter, and  $\text{trunc} < 1$  eliminates these, often increasing power. See Hsu et al. (2013) for comparisons.

## Value

Returns the one P-value for the truncated product. For  $\text{trunc} < 1$ , the distribution is not continuous, but rather attaches positive probability to a P-value of 1.

## Note

Example `mtm` illustrates approximate evidence factors and almost reproduces parts of section 6 of Rosenbaum (2011) using the `mtm` data from Meibian et al. (2008). The second P-value bound differs slightly from that reported in the paper. The reason is that the paper used one scale factor, the median absolute pair difference, for both tests based on all pairs, whereas `senmv` applied to the last two columns of `mtm` uses a scale factor derived just from these two columns. There is no obvious reason to prefer one approach over the other, and the approach in the paper would require a separate R function. Note that in this example,  $y$  is `-mtm` because we test that the controls are low rather than high.

## Author(s)

Paul R. Rosenbaum

## References

Hsu, J. Y., Small, D. S. and Rosenbaum, P. R. (2013) <doi:10.1080/01621459.2012.742018> Effect modification and design sensitivity in observational studies. *Journal of the American Statistical Association*, 108, 135-148.

Meibian, Z., Zhijian, C., Qing, C. et al. (2008) Investigating DNA damage in tannery workers occupationally exposed to tivalent chromium using the comet assay. *Mutation Research* 654, 45-51.

Rosenbaum, P. R. (2010) <doi:10.1093/biomet/asq019> Evidence factors in observational studies. *Biometrika*, 97, 333-345.

Rosenbaum, P. R. (2011) <doi:10.1198/jasa.2011.tm10422> Some approximate evidence factors in observational studies. *Journal of the American Statistical Association*, 106, 285-295. <doi:10.1198/jasa.2011.tm10422>

Rosenbaum, P. R. (2015). Two R packages for sensitivity analysis in observational studies. *Observational Studies*, 1(1), 1-17. Free on-line at [obsstudies.org](http://obsstudies.org)

Rosenbaum, P. R. (2021) <doi:10.1201/9781003039648> *Replication and Evidence Factors in Observational Studies*. Chapman and Hall/CRC.

Zaykin, D. V., Zhivotovsky, L. A., Westfall, P. H. and Weir, B. S. (2002) Truncated product method of combining P-values. *Genetic Epidemiology*, 22, 170-185. <doi:10.1002/gepi.0042>

Zhang, K., Small, D. S., Lorch, S., Srinivas, S. and Rosenbaum, P. R. (2011) Using split samples and evidence factors in an observational study of neonatal outcomes. *Journal of the American Statistical Association*, 106, 511-524.

### Examples

```
# Evidence factor example: see note above.
data(mtm)
senmv(-mtm, gamma=11.7, trim=1)
senmv(-mtm[, 2:3], gamma=2.1, trim=1)
senmv(-mtm, gamma=12, trim=1)
senmv(-mtm[, 2:3], gamma=3, trim=1)
truncatedPbg(c(0.05167572, 0.1527849), trunc=1)
truncatedPbg(c(0.05167572, 0.1527849), trunc=.2)
```



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